

Preview fixation duration modulates identical and semantic preview benefit in Chinese reading

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Abstract Semantic preview benefit from parafoveal words is critical for proposals of distributed lexical processing during reading. Semantic preview benefit has been demonstrated for Chinese reading with the boundary paradigm in which unrelated or semantically related previews of a target word $N + 1$ are replaced by the target word once the eyes cross an invisible boundary located after word N (Yan et al., 2009); for the target word in position $N + 2$, only identical compared to unrelated-word previews led to shorter fixation times on the target word (Yan et al., in press). A reanalysis of these data reveals that identical and semantic preview benefits depend on preview duration (i.e., the fixation duration on the preboundary word). Identical preview benefit from word $N + 1$ increased with preview duration. The identical preview benefit was also significant for $N + 2$, but did not significantly interact with preview duration. The previously reported semantic preview benefit from word $N + 1$ was mainly due to single- or first-fixation durations following short previews. We discuss implications for notions of serial attention shifts and parallel distributed processing of words during reading.

Keywords Eye movement · Parafoveal processing · Semantic · Chinese

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Introduction

The inspection time of each word during sentence reading strongly depends on the words' properties such as its length or frequency of occurrence in a certain language, suggesting that the duration of fixating a word reflects the time needed to process and lexically access its entry in the mental lexicon (for a review see Ranner, 2009). The spatial extent of visual processing during a fixation goes much further beyond the currently fixated word, extending at most up to 4 letters to the left and 14–15 letters to the right of fixation during reading of alphabetic languages (McConkie & Ranner, 1975; Ranner & Bertera, 1979) and 1 character to the left and 2–3 characters to the right of fixation during reading Chinese (C.-H. Tsai & McConkie, 1995; Inhoff & Liu, 1997, 1998). This area, which must be visible for a normal reading rate, is called the *perceptual span* (McConkie & Ranner, 1975). In principle, with a sufficiently short word to the right of a fixated word N , chances are that even the word beyond the next one (i.e., word $N + 2$) may fall into the perceptual span. Whether information of word $N + 2$ can be extracted during reading of alphabetic languages is currently a highly controversial discussion (Ranner, Juhasz, & Brown, 2007; Angele, Slatter, Yang, Kliegl, & Ranner, 2008; for positive results see Kliegl, Risse, & Laubrock, 2007; Risse, Engbert, & Kliegl, 2008; Risse & Kliegl, in press). Here we report a reanalysis of published data and demonstrate that semantic information extraction from words $N + 1$ and $N + 2$ during Chinese reading depends on pre-fixation duration. Pre-fixation duration of words $N + 1$ or $N + 2$ is defined as the time that the reader looks at word N before moving to words $N + 1$ or $N + 2$.

is that phonology plays an important mediating role leading to a word's meaning being activated relatively late (Van Orden, 1987; Van Orden, Pennington, & Stone, 1990; Coltheart, Rastle, Perry, Langdon, & Ziegler, 2001), when compared to some other languages such as Chinese, which we will elaborate below.

In contrast to the role of late semantic activation, evidence from fast priming studies suggests that semantic information facilitates recognition of a foveal target word during a narrow time window at a very early stage with prime durations of about 30 ms (Sereno & Rayner, 1992). In a recent study using a combination of the fast priming and boundary paradigm, Hohenstein, Laubrock, and Kliegl (2010) extended this research to investigate parafoveal semantic priming. Hohenstein et al. varied the duration of parafoveal semantic primes for word $N + 1$. They obtained a semantic priming benefit with a parafoveal prime duration of 125 ms, but not for shorter ones (Experiment 1 and 2). When the saliency of the parafoveal prime word was increased, the semantic priming benefit was significant with an 80-ms but not with the 125-ms parafoveal prime duration (Experiment 3). Thus, in addition to providing evidence for parafoveal processing semantic information in alphabetic languages, the results suggest that semantic priming benefit is time dependent with facilitation due to semantic relatedness of parafoveal priming only during a specific, possibly only early, time window.

Chinese script and semantic priming benefit

Eye-movement control during reading Chinese shares many basic characteristics of alphabetic writing systems (Yan, Kliegl, Richter, Nuthmann, & Shu, 2010). However, there are also important differences, especially with respect to parafoveal processing of semantic information. Chinese script uses square-shaped characters with different levels of visual complexity as the basic writing units; they all occupy the same amount of horizontal extent. There are two important features that make Chinese script particularly well-suited for the demonstration of parafoveal semantic processing. First, in comparison with alphabetic languages, it is generally accepted that Chinese characters are mapped more closely to meaning than to phonology (see Hoosain, 1991, for a summary) whereas the contribution of phonological activation during identification is comparatively small (see Feng, Miller, Shu, & Zhang, 2001, for a review). Second, most Chinese words are only one or two characters long (Yu et al., 1985). Given that a Chinese character typically occupies the space of 3 letters in alphabetic languages (i.e., J. L. Tsai & McConkie, 2003), on average, word $N + 1$ is closer to the point of fixation on word N in Chinese than in alphabetic languages.

Yan, Richter, Shu, and Kliegl (2009) investigated parafoveal processing of Chinese words in position $N + 1$ by manipulating the first character in two-character words and found a reliable priming benefit for characters semantically related to the target. Yan, Kliegl, Shu, Pan, and Zhou (in press) used the same material as Yan et al., but they moved the target word to the $N + 2$ position by inserting a high-frequency or low-frequency word in position $N + 1$. They reported an $N + 2$ priming benefit (see also Yang, Wang, Xu, & Rayner, 2009), but only when the priming character was identical to the target character; the effect was larger with a high-frequency word $N + 1$. Importantly, there was no significant priming benefit

When the pre-identified character is semantically related to the target but there is a trend in this direction.

In the boundary paradigm the pre-identifications are always either available or denied for the entire duration of the fixation prior to the boundary. This raises the possibility that the failure to find a reliable semantic pre-identification benefit for word $N + 2$ in Chinese could be due to a dependence of this effect on a specific time window. Hohenstein et al.'s (2010) results suggest that parafoveal semantic information may facilitate processing as early as 125 ms. As preboundary fixation durations are usually much longer than 125 ms (and under the readers' control), these long pre-identifications may have masked a semantic pre-identification benefit during a specific time window in alphabetic scripts.

Sequential attention shifts, processing gradient models

Statistically reliable evidence for semantic preprocessing of word $N + 2$ provides important constraints for theoretical accounts of serial attention shift (e.g., Engbert & Kliegl, 2001; Reichle, Li, Ersdege, Pollatsek, & Raner, 2009) and processing gradient models of eye-movement control (e.g., Engbert & Kliegl, 2010; Engbert, Nuthmann, Richter, & Kliegl, 2005; Reill & Radach, 2003, 2006). Serial attention shift (SAS) models like E-Z Reader (Reichle, Pollatsek, Fisher, & Raner, 1998; Reichle et al., 2009; see Engbert & Kliegl, 2001, for a different variant) assume that lexical processing occurs only at the attended word and that attention shifts to the next word only after lexical access is completed. Thus, semantic pre-identification is problematic for serial attention shift models (e.g., Reichle et al., 2009, for a review). On the other hand, processing gradient (PG) models such as SWIFT (Engbert et al., 2005) or Glenmore (Reill & Radach, 2003, 2006) assume distributed lexical processing in the perceptual span. As a consequence of this principle, PG models generally allow semantic preprocessing for words $N + 1$ and even for words $N + 2$ as long as they are in the perceptual span. However, due to the acuity-related decrease of processing efficiency with eccentricity from the current fixation location, semantic preprocessing for word $N + 2$ might be too weak to be detected in alphabetic languages. In Chinese, as a language in which the information is more densely packed, the semantic information extraction has been shown for word $N + 1$ and may even be visible for word $N + 2$ for an appropriate time window of pre-identification.

The present study

The current study reports a reanalysis of Yan et al. (2009) and Yan et al. (in press). In the boundary paradigm, pre-identification duration of word $N + 1$ and word $N + 2$ is controlled by participants' fixations on the preboundary word N . When a saccade is executed that crosses the invisible boundary, the display change of word $N + 1$ and word $N + 2$ is triggered and terminates the parafoveal prime. Therefore, the variability of pre-identification durations may act like different parafoveal prime durations and can be used as a covariate for the size and direction of various pre-identification effects.

Using the data from the two prior studies, we focus on the size of various information processing effects (i.e., identical, semantic, orthographic, and phonological relative to unrelated pre-lexical words) as a function of the pre-lexical duration for parafoveal processing of word $N + 1$ and $N + 2$ in Chinese reading.

Method

Subjects

All participants of the experiments were native Chinese students from Beijing Normal University with normal or corrected to normal vision. For the eye-tracking experiments, 48 students were tested for Data Set 1 with a manipulation of word $N + 1$ (Yan et al., 2009) and an independent sample of 74 students contributed to Data Set 2 with a manipulation of word $N + 2$ (Yan et al., in press). Also, 51 students who did not participate in the two eye-tracking experiments were recruited for three norming studies of relatedness between pre-lexical words and targets.

Material

Forty-eight simple non-compound characters were selected as targets, which served as the first character of word $N + 1$ in Data Set 1 and the first character of word $N + 2$ in Data Set 2. Each target character was embedded into a two-character target word, only the identical condition provided legal word-level pre-lexical words. For each target character, four types of pre-lexical characters were selected for orthographically related, phonologically related, semantically related, and unrelated pre-lexical conditions. As shown in Table 1, there were no differences between the pre-lexical types with respect to visual complexity as indexed by number of strokes

Table 1 Means (standard deviations) of frequency (per million), number of strokes and relatedness ratings of target and pre-lexical characters

	Target	Pre-lexical				
		Identical	Orthographic	Phonological	Semantic	Unrelated
Example	永	永	水	用	久	向
Meaning	Forever	Forever	Water	Usage	Long	Towards
Pronunciation	ong3	ong3	shui3	ong4	jiu3	iang4
Frequency	1,150 (1,728)	1,150 (1,728)	1,154 (1,435)	1,197 (1,757)	1,164 (1,721)	1,163 (1,573)
No. of strokes	5.0 (2.1)	5.0 (2.1)	4.8 (1.8)	5.1 (1.9)	5.5 (2.6)	4.9 (1.9)
Orthographic rating			3.8 (0.8)	1.6 (0.3)	1.5 (0.4)	1.6 (0.3)
Phonological rating			1.2 (0.3)	4.3 (0.6)	1.2 (0.2)	1.1 (0.2)
Semantic rating			1.2 (0.3)	1.2 (0.1)	4.1 (0.6)	1.2 (0.2)

Ratings set in bold signify that independent ratings matched the intended experimental manipulation

($F = 1.0$, $p > .1$) and frequency ($F < 1$). The three relatedness ratings nicely reflected the intended design. Due to non-significant phonological pre-emption for word $N + 1$ in the first attention analysis reported in Yan et al. (2009), this condition was removed from Data Set 2.

The invisible boundary that triggered the display change was located just to the left of character $N + 1$, which is the first character of the target word (word $N + 1$) in Data Set 1, and a single-character word prior to the first character of the target word (word $N + 2$) in Data Set 2. Eye movements were recorded with an EyeLink II system (500 Hz). Single sentences were presented on the vertical position one-third from the top of the screen of a 19-inch ViewSonic G90f monitor (1,024 × 768 resolution; frame rate 100 Hz) for Data Set 1 and a 21-inch Dell Trinitron Monitor (1,280 × 1,024 resolution; frame rate 100 Hz) for Data Set 2. Therefore, it took at most 16 ms to complete the display change for both data sets. The words before the boundary (i.e., word N) were also all single-character words. Each sentence was only presented once to a participant with the different pre-emption. A set of example sentences is shown in Fig. 1. Full details concerning the material, apparatus and procedure are available in Yan et al. (2009) and Yan et al. (in press).

(A)

Identical preview:

Orthographical preview:

Phonological preview:

Semantic preview:

Unrelated preview:

Target sentence:

Data analysis

Data were reduced to a certain format using an algorithm for the binocular detection of saccades (Engbert & Kliegl, 2003). Sentences containing a blink or loss of measurement were deleted (i.e., 18% in Data Set 1 and 5% in Data Set 2). Analyses were based on right-eye saccades during first-pass reading. We distinguish between first saccade durations (FFDs; the first saccade on a word, irrespective of the number of saccades), single saccade durations (SFDs; cases in which a word was inspected with one saccade), and gaze durations (GDs; the sum of saccades during the first reading of the word). Cases with FFDs shorter than 60 ms or longer than 600 ms were excluded (1% of all saccades in Data Set 1 and 2% in Data Set 2). Further, trials with regressions from word N or $N + 1$ for Data Set 1 as well as from word N , $N + 1$ or $N + 2$ for Data Set 2 were excluded (10% trials in Data Set 1 and 11% trials in Data Set 2). For the preprime conditions in Data Set 1, there were 1,052 observations in the LMM model for FFD and GD analyses and 769 observations for SFD analyses; for the four preprime conditions in Data Set 2, there were 4,024 observations in the LMM model for FFD and GD analyses and 3,385 observations for SFD analyses.

Inferential statistics are based on planned comparisons for the related and the identity primes with the unrelated prime as reference. Estimates are based on a linear mixed model (LMM) with crossed random effects for subjects and items using the *lmer* program of the *lme4* package (Bates & Maechler, 2010) in the R environment for statistical computing and graphics (R-Core Development Team, 2010). Estimates larger than 2 SE (i.e., $t > 2$) are interpreted as significant. Analyses of residuals and inspection of duration distributions strongly suggested that log-transformation is required to meet LMM assumptions. Therefore, we used log-transformed durations for LMMs.

Results

The main goal of the present research was to test whether the duration of the saccade prior to the display change in cases when only a single saccade is made on the preboundary word modulates prime benefit from semantically related parafoveal words in positions $N + 1$ and $N + 2$ in reading of Chinese. FFDs, SFDs, and GDs on word $N + 1$ (Data Set 1) and $N + 2$ (Data Set 2) were used as dependent variables. In the LMMs, these effects of interest translate into interactions between the continuous predictor of single-saccade duration on preboundary word N and planned comparisons of semantic, orthographic, phonological (only in Data Set 1), and identical primes with an unrelated prime as baseline (i.e., treatment contrasts with unrelated prime as reference category). Main effects were evaluated at the mean of the log prime SFD (i.e., the covariate was centered). Thus, the intercept represents the mean log FFD, mean log SFD, or mean log GD on the target word for the unrelated condition. Analyses using prime SFD as covariate yielded the clearest dissociation of effects, possibly because single-saccade cases can be mislocated saccades (Nuthmann, Engbert, & Kliegl, 2005) and are reliable indicators of successful parafoveal word segmentation (Yan et al., 2010).

We note that similar trends (not all statistically significant) were also present for FFDs and GDs on the preboundary order N . We also test the critical results in post-hoc comparisons for short and long pre-ice durations, using the mean of the log-transformed single ablation duration as cut-off point.

Pre-ice benefit for order $N + 1$

shown for FFDs (panel A) and GDs (panel B). The vertical line indicates the mean log preview duration (i.e., the value at which main effects are evaluated).

With FFD as dependent variable (Fig. 2a), the identical preview effect was remarkably large and also largely independent of preview duration (i.e., distance between identical and unrelated conditions is large and the hypothesis that the two lines are parallel cannot be rejected; interaction t -values for FFDs, also SFD, < 1).

On 17, 30, 25, 34 and 29% (for identical, orthographic, phonological, semantic and unrelated preview conditions, respectively) of all valid trials, first fixations on target words followed by rejections. With GD as dependent variable (Fig. 2b), the identical preview benefit significantly increased with preview duration ($b = -0.28$, $SE = 0.14$, $t = -2.0$, for the interaction of identical vs. unrelated preview and preview duration).¹ The increase in the preview benefit resulted from the divergence in GD for unrelated and identical previews; neither the numeric GD increase for the unrelated preview ($t = 1.36$) nor the numeric GD decrease for the identical preview ($t = -1.37$) was significant by itself.

The similarity between FFDs and GDs in slopes for the identical conditions (bold dotted lines in Fig. 2) suggests that rejection rate did not depend on preview duration. The divergence in slopes for the unrelated conditions (negative for FFDs and positive for GDs; bold solid lines in Fig. 2) suggests that rejection rate increased during preview. This was confirmed in post-hoc analyses of rejection rate, using a binary measure of hitting the target once or more than once as dependent variable in a generalized linear mixed model (GLMM): Rejection rate increased significantly during preview in the unrelated preview condition ($b = 0.25$, $SE = 0.10$, $t = 2.5$), but rejection rate did not decrease significantly in the identical preview condition ($b = -0.07$, $SE = 0.11$, $t = -0.6$). Traditionally, the unrelated preview condition serves as the baseline for the computation of the preview benefit. The increase in rejection rate with preview duration in this condition may be interpreted as evidence for a preview cost. This is a very important result because it suggests that classical preview benefits may arise in part as a consequence of *preview cost* associated with long previews of unrelated words.

Semantic preview benefit

The main effect of semantic preview was also significant for FFDs and marginally significant for SFDs ($b = -0.07$, $SE = 0.03$, $t = -2.7$, and $b = 0.06$, $SE = 0.03$, $t = -1.8$, respectively) and there was a numeric trend for GDs ($b = -0.06$, $SE = 0.04$, $t = -1.6$). These (tendencies to) main effects were strongly qualified by interactions with preview duration (i.e., preview SFDs; $b = 0.16$, $SE = 0.10$, $t = 1.7$, and $b = 0.23$, $SE = 0.12$, $t = 2.0$, for FFDs and SFDs, respectively; see Footnote 1). The dashed bold line (semantic preview) and the solid bold line

¹ We also tested the interaction between preview duration and preview benefits in a LMM with subgroup as two-level factor replacing the covariate (i.e., logarithm preview single fixation duration), which is more compatible with the traditional ANOVA route. In this analysis we failed to replicate the significant interactions (identical preview benefit in GD analysis: $b = -0.09$, $SE = 0.07$, $t = -1.2$; semantic preview benefit: $b = 0.08$, $SE = 0.05$, $t = 1.5$; $b = 0.09$, $SE = 0.06$, $t = 1.5$; for FFD and SFD analyses, respectively). We present this also as evidence that not everything is significant in LMM (as is sometimes surmised).

(unrelated pre-lexical) in Fig. 2a shows that FFDs with semantic pre-lexical were as short as those for identical pre-lexical given a 150 ms pre-lexical duration, but were as long as those for unrelated pre-lexical with a pre-lexical duration of 400 ms. Thus, the semantic pre-lexical benefit differed from the identical pre-lexical benefit: the semantic pre-lexical benefit was large for short pre-lexical and vanished with increasing pre-lexical duration whereas identical pre-lexical benefit was present for all pre-lexical durations. The effects were not significant with GD as the dependent variable.

Orthographic and phonological preview benefits

There was a significant main effect for orthographic pre-lexical in FFDs ($b = -0.06$, $SE = 0.03$, $t = -2.1$; SFDs: $t = -1.7$, GDs: $t = -1.7$). There was no significant effect of the phonological pre-lexical condition (all t -values < 1.4).

None of the interactions between orthographic or phonological pre-lexical and pre-lexical duration was significant, that is the slopes for the orthographic (dot-dash) and phonological (dashed) pre-lexicals did not differ significantly from the one for the unrelated-pre-lexical baseline. In a follow-up LMM with identical pre-lexical as reference conditions, these slopes were not significantly different from this condition either (all t -values < 1.4).

Preview effects for grouped short and long previews

Identical preview. As a further illustration of the significant interaction, we separated trials into two subgroups with a cutoff point of mean log pre-lexical SFD of 217 ms; the value at which the main effects in the above LMMs were evaluated (see Table 2). This criterion led to 572 observations for FFD and GD analyses as well as 420 observations for SFD analysis for the short pre-lexical group, and 480 observations for FFD and GD analyses as well as 349 observations for SFD analysis for the long pre-lexical group. Results indicated that identical pre-lexical benefit in GDs was significant in each of the subgroups and increased in effect size with

Table 2 Means (standard errors) of first-mention duration (FFD), single-mention duration (SFD) and gaze duration (GD) on word $N + 1$ from Data Set 1, broken down by mean log pre-lexical single-mention durations

	No. obs	Identical	Orthographic	Phonological	Semantic	Unrelated
<i>Short preview</i>						
FFD	572	216 (9)	235 (10)	249 (9)	231 (9)	259 (7)
SFD	420	214 (10)	240 (12)	256 (11)	237 (11)	263 (8)
GD	572	260 (16)	307 (17)	322 (17)	308 (16)	332 (15)
<i>Long preview</i>						
FFD	480	219 (12)	255 (11)	254 (11)	254 (12)	262 (9)
SFD	349	218 (13)	246 (13)	257 (13)	259 (14)	259 (11)
GD	480	257 (23)	333 (23)	326 (23)	340 (23)	370 (20)

Means and standard deviations are computed across grand means

increased pre-lexical duration ($b = 0.23$, $SE = 0.05$, $t = 4.7$ and $b = 0.31$, $SE = 0.06$, $t = 5.2$ for short and long pre-lexical, respectively).

Semantic preview. The semantic preview benefit was also modulated by pre-lexical duration: It was significant for short pre-lexical ($b = 0.10$, $SE = 0.03$, $t = 2.9$ and $b = 0.10$, $SE = 0.04$, $t = 2.3$; for FFD and SFD analyses, respectively) but not for long ones (both t -values < 0.7).

Orthographic and phonological preview. Orthographic preview benefit was significant for short pre-lexical, $b = 0.09$, $SE = 0.04$, $t = 2.3$ and $b = 0.09$, $SE = 0.04$, $t = 2.0$; for FFD and SFD analyses, respectively; both t -values were smaller than 1 for long pre-lexical. Note in the LMM this effect was significant as a main effect. The phonological preview benefit tended to be significant for long pre-lexical ($b = 0.10$, $SE = 0.06$, $t = 1.7$ and $b = 45$ ms, $SE = 23$ ms, $t = 1.9$; for analyses in log-transformed and original metrics, respectively) compared to short pre-lexical (both t -values < 0.7).

Pre-lexical benefit for word $N + 2$

Skipping of word $N + 1$

In the second data set with target words in position $N + 2$, all of the words in position $N + 1$ were one character long. Consequently, there was a high skipping probability associated with this word (54%). It is well known that durations after skipped words are longer than on average. In the present experiment skipping of $N + 1$ increased FFDs by 23 ms, SFDs by 23 ms and GDs by 58 ms on target word $N + 2$. These effects were highly significant in the LMM ($b = 0.09$, $SE = 0.01$, $t = 7.7$; $b = 0.08$, $SE = 0.01$, $t = 8.2$; $b = 0.18$, $SE = 0.01$, $t = 12.5$; for FFD, SFD and GD analyses, respectively). Skipping of word $N + 1$, however, did not interact with pre-lexical duration or pre-lexical conditions (all t -values < 1.4).

Identical preview

In general, pre-lexical effects were much weaker for word $N + 2$ than for word $N + 1$ (compare Fig. 3 for word $N + 2$ and Fig. 2 for $N + 1$). The largest effect in this data set was again the main effect of identical preview ($b = -0.05$, $SE = 0.01$, $t = -3.8$; $b = -0.05$, $SE = 0.01$, $t = -3.7$; $b = -0.07$, $SE = 0.02$, $t = -4.2$); for FFD, SFD and GD analyses, respectively, but none of the interactions with pre-lexical duration were significant (all $t < 1$). In a follow-up LMM, FFDs, SFDs, and GDs after identical preview were also significantly shorter compared to semantic or orthographic preview (all $t > 2.65$).

Semantic preview

Despite the strongly reduced pre-lexical modulation, the bold dashed line for semantic preview and bold solid line for unrelated preview represent a significant interaction

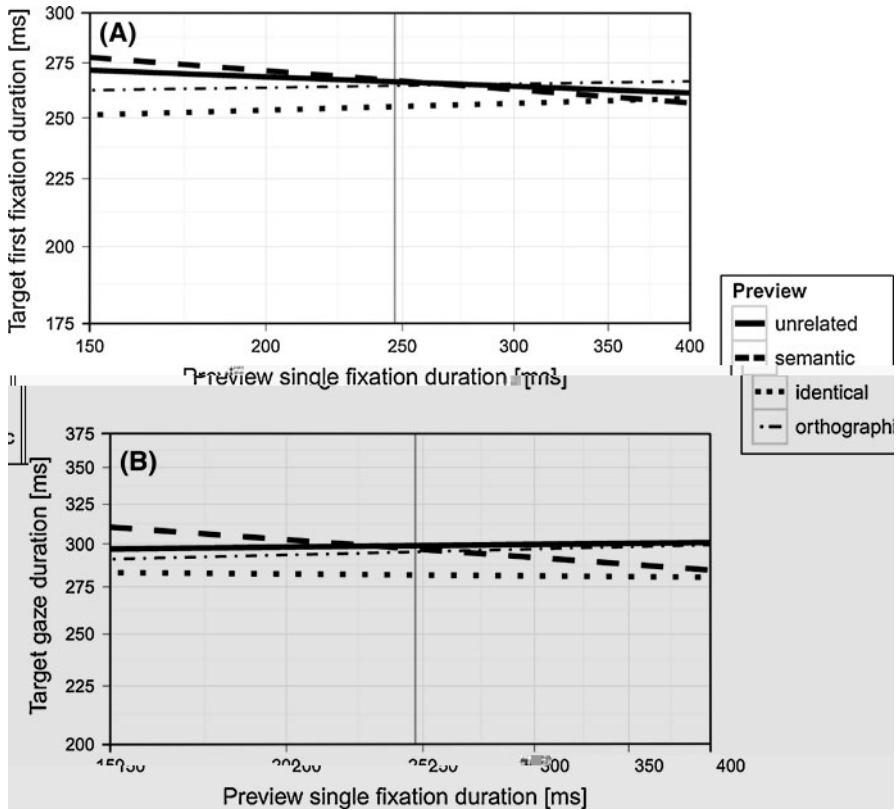


Fig. 3 Linear regression of first fixation duration (a) and gaze duration (b) on word $N + 2$ on single fixation duration on word N for unrelated (*bold-solid*), semantic (*bold-dashed*), identical (*bold-dotted*) and orthographic (*simple dashed*) preview conditions using logarithmic scales for both axes. The vertical line indicates the mean log single-fixation duration on word N . Between-subject and between-item differences for dependent variable and covariance in the LMM were removed prior to regressions. Figure as generated with *ggplot2* (Wickham, 2009)

between preview duration and semantic preview benefit for GD analysis ($b = -0.11$, $SE = 0.05$, $t = -2.0$).² Again, we observed a crossover pattern, but this time the semantic preview effect was negative for shorter than average preview durations and positive for longer than average ones (i.e., semantic preview benefit; see Fig. 3b). This negative difference was not significant in the subgroup of short previews, but neither was the positive difference for long previews (see below). The interaction was not predicted. Therefore, the result is in need of independent replication before it is used for substantive interpretations.

² This interaction also reached significance in a LMM using a logic grouping factor ($b = -0.06$, $SE = 0.03$, $t = -1.9$).

Table 3 Means (standard errors) of first-mention duration (FFD), single-mention duration (SFD) and gaze duration (GD) on word $N + 2$ from Data Set 2, broken down by mean log pre-mention single-mention durations

Means and standard deviations are computed across grand means

	No. obs	Identical	Orthographic	Semantic	Unrelated
<i>Short preview</i>					
FFD	2,222	248 (5)	257 (4)	266 (5)	263 (6)
SFD	1,855	245 (5)	259 (5)	265 (5)	260 (6)
GD	2,222	267 (8)	281 (8)	293 (8)	286 (9)
<i>Long preview</i>					
FFD	1,802	259 (6)	275 (6)	271 (6)	270 (6)
SFD	1,530	259 (6)	274 (6)	273 (6)	272 (6)
GD	1,802	285 (9)	298 (9)	298 (9)	312 (9)

Orthographic preview

The main effect of orthographic preview and its interaction with preview duration were not significant (all t -values < 1.2).

Preview benefit for grouped short and long previews

Post-hoc breakdown of trials by mean log preview duration (i.e., 247 ms in original metric) did not reveal significant semantic or orthographic preview benefits for any of the groups.³ As shown in Table 3, there was only a numerical trend of a semantic and orthographic preview benefit with long previews for GD analyses (semantic preview benefit: $b = 0.04$, $SE = 0.02$, $t = 1.6$; $b = 0.02$, $SE = 0.02$, $t = 0.9$; for trials with long and short previews, respectively). For analyses of orthographic preview benefit: $b = 0.03$, $SE = 0.02$, $t = 1.4$; $b = 0.02$, $SE = 0.02$, $t = 0.7$; for trials with long and short previews, respectively). The lack of significance in the post-hoc analysis is a consequence of the loss of statistical power associated with using a dichotomized factor derived from a continuous covariate of preview duration (e.g., Baayen, 2008).

Finally, the identical preview benefit was significant for both groups ($M = 27$ ms, $b = 0.08$, $SE = 0.02$, $t = 3.2$ and $M = 19$ ms, $b = 0.06$, $SE = 0.02$, $t = 2.7$; for trials with long and short previews, respectively), with a numerically larger effect for long previews.

³ Split of trials by preview single-mention duration at 240 ms led to balanced groups in number of observations and more representative demonstrative results: For semantic preview benefit with long previews, No. obs = 1,958, $b = 0.04$, $SE = 0.02$, $t = 1.8$ and $b = 17$ ms, $SE = 9$, $t = 2.0$; for analyses in log-transformed and original metrics, respectively; for orthographic preview benefit with long previews, $b = 0.03$, $SE = 0.02$, $t = 1.5$ and $b = 16$ ms, $SE = 9$, $t = 1.8$; for analyses in log-transformed and original metrics, respectively. Identical preview benefit was also numerically larger for long ($b = 0.08$, $SE = 0.02$, $t = 3.5$ and $b = 30$ ms, $SE = 9$, $t = 3.4$; for analyses in log-transformed and original metrics, respectively) than for short previews ($b = 0.06$, $SE = 0.02$, $t = 2.7$ and $b = 19$ ms, $SE = 8$, $t = 2.4$; for analyses in log-transformed and original metrics, respectively). All other t -values were smaller than 1.

Discussion

In alphabetic languages, studies using the boundary paradigm in a natural sentence reading task have so far failed to demonstrate perceptual benefits of semantically related parafoveal prelines (see Raner et al., [2003](#), for a review). Recently, such

Liersch (2005) who partitioned their data on the median for participants and conditions could be due to reduced statistical power for dichotomized covariates.

The detection of the increase of pre-lexical benefit from word $N + 1$ across pre-lexical duration reported here probably requires the use of a continuous measure of pre-lexical duration and statistically more powerful techniques (such as LMM) than were used in the earlier studies. Our results that identical pre-lexical benefit increased with pre-lexical duration is in agreement with Schroens, Vitu, Brsbaert, and d'Ydewalle (1999) who presented a sequence of three words with an invisible boundary between the first and second word of the triad, manipulating pre-lexical offset of word $N + 1$ during preboundary transitions on word N , and reported larger pre-lexical benefit on target word $N + 1$ with increasing pretarget durations.

Semantic pre-lexical benefit

The 'orthograph-to-phonology-to-semantics' route (Coltheart et al., 2001; Van Orden, 1987; Van Orden et al., 1990) assumes a sequential activation with access to word meaning in a relatively late stage. We suspect that this route to word recognition is more dominant in English than Chinese, which is known as a writing system with a close association between graphic form and meaning (see Hoosain, 1991, for a summary). For example, there is strong evidence for direct access from orthograph to semantics with phonological mediation bypassed under some circumstances (Chen & Shu, 2001; Meng, Jian, Shu, Tian, & Zhou, 2008; Zhou & Marslen-Wilson, 1999, 2000). Against this background of research, it is not surprising that parafoveal pre-lexical processing of a semantically related character significantly reduces the subsequent fixation time on this target. Indeed, recent studies of Chinese reading using the boundary paradigm demonstrated reliable semantic pre-lexical benefit from word $N + 1$ for simple (Yan et al., 2009) and compound characters (Yang, Wang, Tong, & Raner, 2010).

The failure to find evidence for a semantic pre-lexical benefit for word $N + 1$ in alphabetic scripts (Altarriba et al., 2001; Raner et al., 1986) and for word $N + 2$ in Chinese (Yan et al., in press) may have been due to the fact that the pre-lexical mask was displayed throughout the whole fixation duration on word N . Hohenstein et al. (2010) controlled the duration of the parafoveal semantic pre-lexical and demonstrated that semantic pre-lexical benefit might be restricted to an early time window.

The present study is a reanalysis of data from Yan et al. (2009) and Yan et al. (in press). The dependence of semantic pre-lexical benefit on pre-lexical duration was tested with the interaction between pre-lexical duration and the contrast of semantic and unrelated pre-lexical. Facilitation due to semantic pre-lexical offset of word $N + 1$ was observed only for pre-lexical durations shorter than 217 ms; semantic pre-lexical benefit was not significant for long pre-lexical durations. The LMM results suggest that accumulation of information specific to the meaning of the semantically related pre-lexical word may interfere with lexical access of the target word. FFDs following a short semantic pre-lexical are similar to those following a short identical pre-lexical, but FFDs following a long semantic pre-lexical are similar to those following a long

unrelated pre-lexical. Thus, a semantically related pre-lexical word changes from being functionally identical with the target word itself to being functionally unrelated to the target word.

Orthographic and phonological pre-lexical benefit

A time dependence analogous to the one observed for semantic pre-lexical benefit was also obtained for orthographic pre-lexical benefit with significant facilitation for short pre-lexical words $N + 1$. Finally, in line with a relatively late stage of phonological activation in Chinese sentence reading (Feng et al., 2001), the phonological pre-lexical benefit for words $N + 1$ was mainly observed in trials with long pre-lexical words.

Time course of parafoveal processing and attention

The time course of parafoveal processing has been discussed as an opportunity to test different assumptions regarding the timing of parafoveal processing.

We are much doubt that any of the currently available computational models, such as the E-Z Reader SAS model (Reichle et al., 1998; 2009; see Raner, Li, & Pollatsek, 2007, for an adaptation for reading Chinese) or models built around the assumption of processing gradients, such as the SWIFT model (Engbert et al., 2005) or Glenmore (Reilly & Radach, 2003, 2006) are read to reproduce such competition of lexical activations. With its well-defined linguistic processing components, the Glenmore model might have the best chance to capture the time-dependent inhibition effects of parafoveal extracted incorrect information.

In general, the present results favor the notion of a 'spot' in time at which parafoveal information is integrated across saccades (e.g., Schiepers, 1980). They are certainly compatible with the assumption of parallel distributed processing. Statistically reliable evidence for semantic information extraction (either facilitation or inhibition) from word $N + 1$ is in favor of parallel models.

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